

1 The impact of agricultural landscape diversification on U.S. crop production

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16
17 **Abstract:** The last century has seen a dramatic simplification of global landscapes, driven largely by
18 the expansion and intensification of agriculture. Landscape simplification has known negative
19 impacts on ecosystem health and function; however, less is known about how landscape
20 simplification affects agricultural production. There is mounting *field-scale* evidence that simplification
21 can reduce agricultural production by eroding the ecosystem processes on which agricultural systems
22 depend; however, many of these processes emerge not at the field scale, but from complex
23 interactions between land use, biophysical context, and human activity at the *landscape scale*. This
24 research uses hierarchical Bayesian models to estimate the relationship between landscape-scale
25 agricultural diversity and the yields of corn, soy, and winter wheat in the coterminous United States.
26 We find that the yields of corn and winter wheat increase by as much as 20% in highly diversified
27 agricultural systems. Our findings also indicate that (1) crop production is more responsive to the
28 number of distinct crop types cultivated on a landscape than their cultivated extent and that (2)
29 increasing diversity in agricultural systems that are already diverse brings the highest yield gains. Our
30 models provide strong evidence at national and regional scales that agricultural diversification—an
31 intervention with known ecosystem benefits—can increase crop production.

32 33 **Highlights:**

- 34 • The yields of corn and winter wheat increase by as much as 20% in highly diversified
35 agricultural systems, while soy yields increase by nearly 5%.
- 36 • Crop production is more responsive to the number of agricultural land use categories in a
37 system than the relative cultivated extent of each category.
- 38 • Increasing agricultural diversity in systems that are already diverse brings the highest yield
39 gains.

40
41 **Keywords:** Diversity, crop production, United States

42
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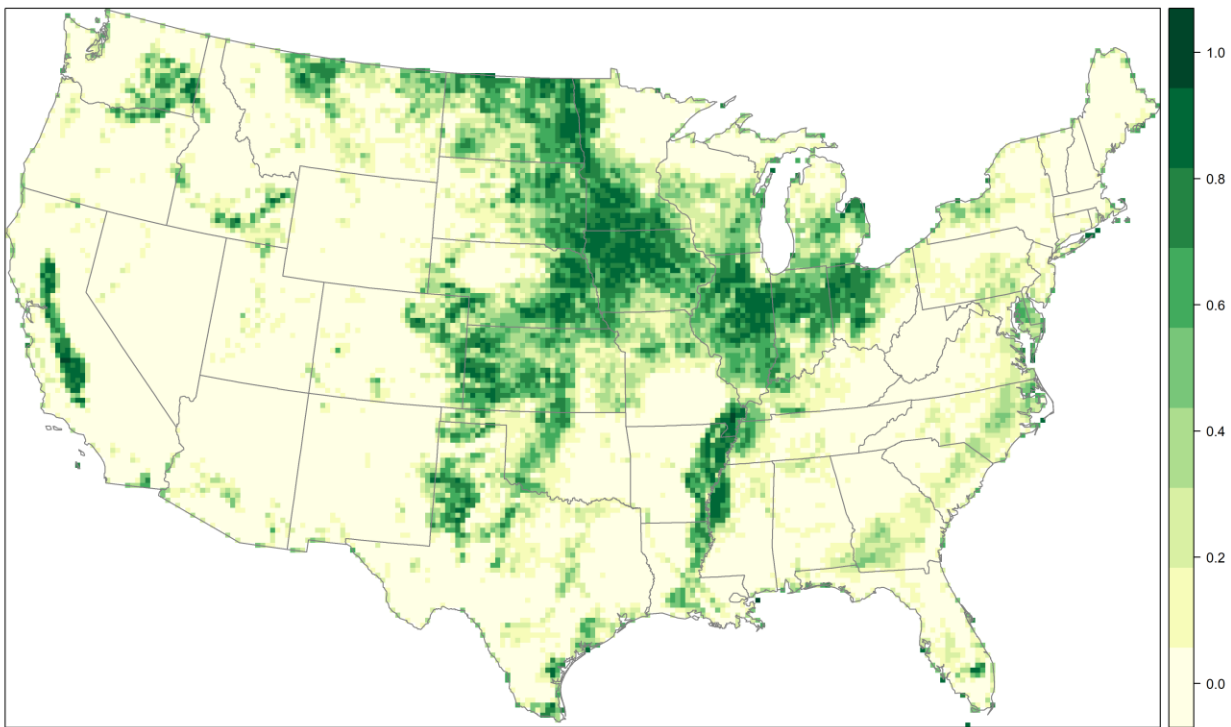
48 **The impact of agricultural landscape diversification on U.S. crop production**

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50 **1.0 INTRODUCTION**

51

52 The last century has seen a dramatic simplification of global landscapes, driven largely by the
53 expansion and intensification of agriculture (Aguilar et al., 2015; Khoury et al., 2016; Landis, 2017).
54 Agriculture now covers one-third of global land, making it the most significant “engineered
55 ecosystem” on the planet (Zhang et al., 2007). In the U.S., agriculture accounts for over 50 percent
56 of total land area (Figure 1) – and over half of this land is cultivated with corn, soy, or wheat
57 (Bigelow & Burchers, 2012). Simplified agricultural landscapes with low levels of natural habitat and
58 plant diversity are optimized for crop production (Meehan et al., 2011, Grab et al. 2018); however,
59 they are also associated with soil degradation, loss of habitat, reductions in water quality, and loss of
60 species diversity (Bommarco, Kleijn, & Potts, 2013; Hendrickx et al., 2007; Landis, 2017; McDaniel,
61 Tiemann, & Grandy, 2014; Tiemann, Grandy, Atkinson, Marin-Spiotta, & McDaniel, 2015;
62 Tschardt et al., 2012). These negative environmental impacts in turn erode the ecosystem
63 processes essential to crop production such as pollination, pest management, water retention, and
64 nutrient supply (Swift et al., 2004; Zhang et al., 2007). This implies that over time agriculturally-
65 driven landscape simplification may diminish agricultural productivity.



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67 **Figure 1:** The proportion of agricultural land use across the U.S. as indicated by the USDA
68 CropScape dataset (2017). Dark green indicates a higher intensity of agricultural land use.

69

70 An axiom of ecology and sustainability science is that diversity increases the health and
71 function of complex systems (Bommarco et al., 2013; Khoury et al., 2016; Walker et al., 2004).
72 Evidence from hundreds of experiments confirms that diversity, in and of itself, is essential to

73 ecosystem productivity (Cardinale et al., 2012, 2006; Hooper et al., 2012; Loreau et al., 2001; Tilman
74 et al., 2014, 2012). Despite being inextricably linked to ecological systems, agricultural systems are
75 often purposefully managed to *reduce* species diversity to increase harvestable yields. Farmers, who
76 play a central role in the selection of which species are present on a landscape, are influenced by
77 policies and institutions endorsing specialization as a tool to increase agricultural productivity
78 (Nassaur, 2010; Roesch-McNally, Arbuckle, & Tyndall, 2018; Yoshida, Flint, & Dolan, 2018).
79 Whether this economic assumption aligns with biological reality is highly contested (Cassman, 1999;
80 Davis et al., 2012; Kremen & Miles, 2012; Reiss & Drinkwater, 2018; Virginia et al., 2018).

81 Field-scale experiments suggest that—as in ecological systems—diversity can actually *increase*
82 agricultural production (Li et al., 2009; Ojha & Dimov, 2017; Smith, Gross, & Robertson, 2008;
83 Tschardt et al., 2005). Smith and colleagues (2008) found that corn yield increases were 100
84 percent higher in diverse agricultural systems as compared to monoculture systems. Pywell et al.
85 (2015) and more recently Schulte et al. (2017) found that transforming even a small percentage of
86 agricultural land to wildlife habitat maintained or improved yields. Several papers have found that
87 crop diversity is associated with reduced yield volatility over time (Abson, Fraser, & Benton, 2013;
88 Di Falco & Perrings, 2005; Weigel, Koellner, Poppenborg, & Bogner, 2018). Research suggests that
89 these yield improvements are driven by the positive impact of diversification on the ecosystem
90 services essential to crop production, including pest management (Bommarco et al., 2013; Chaplin-
91 Kramer et al., 2011; Gardiner et al., 2009), soil health (McDaniel et al., 2014; Tiemann et al., 2015),
92 and pollinator diversity (Schulte et al., 2017; Tschardt et al., 2005).

93 Almost all of the existing evidence linking diversity to increased agricultural production is at
94 the field-scale; however, many of the ecological processes on which agricultural systems depend
95 emerge not at the field-scale, but from complex interactions between land use, biophysical context,
96 and human activity at the *landscape scale*. Landscape composition and configuration have been shown
97 to affect many of the ecosystem services essential to agriculture, such as water quantity and quality,
98 pollination, pest regulation, carbon storage, and climate management (Li et al., 2009; Swinton, Lupi,
99 Robertson, & Hamilton, 2007). In addition, many ecosystem services essential to agriculture such as
100 pollinator movement and water flow are generated far from the agricultural fields that benefit from
101 them. Therefore, field-scale efforts to diversify may be negated by landscape simplification and
102 conversely, landscape-scale diversification may benefit localized monoculture systems (Tschardt et
103 al., 2005). For these reasons and the mounting evidence linking field-scale diversification to
104 increased ecosystem services and agricultural productivity, we hypothesize that landscapes with
105 higher levels of agricultural diversity will support more productive agricultural systems.

106 As agriculture becomes the most widespread use of land on Earth, there is a critical need to
107 determine how and why agriculturally-driven landscape change affects agricultural production. This
108 research uses Bayesian hierarchical modeling to estimate the relationship between agricultural
109 diversity and the yields of corn, soy, and winter wheat in counties across the coterminous United
110 States while controlling for seasonal climate, spatiotemporal dependencies, and regional factors
111 known to influence yield. Our results indicate that agricultural diversity is associated with increased
112 agricultural productivity and that it is primarily the number of agricultural land use categories, rather
113 than their relative cultivated extent, that drive these yield gains. Regional variability in our models,
114 however, highlights the continued importance of local and regional analyses to assess the complex
115 assemblage of socio-ecological factors that mediate the diversity-productivity relationship across
116 space and time.

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121 2.0 METHODS

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123 2.1 Data

124 The county-season is the smallest spatiotemporal unit at which public yield data is available
125 nationally (USDA NASS, 2019); however, land use and weather data are available at much higher
126 spatiotemporal resolutions. To resolve this scalar mismatch and preserve as much information as
127 possible, we constructed county-scale indicators of cumulative seasonal weather exposure from
128 gridded daily temperature and precipitation data and computed three indicators of county-scale
129 agricultural diversity from annual 30 meter land use data. We extracted gridded land use and
130 weather data to the county scale and merged this data with county-level yield estimates for corn, soy,
131 and winter wheat for all counties in the conterminous U.S. (n=3108) from 2010 to 2016. We focus
132 on corn, soy, and winter wheat because of their importance to the global economy and their
133 prevalence on U.S. agricultural landscapes. Since the 1960s, harvested soy and corn acreage has
134 increased by 76 percent (74 million acres), today covering about 90 million and 89 million acres
135 respectively (Bigelow & Borchers, 2017). Wheat – including winter, durum, and spring wheat –
136 comprises the third largest acreage in the U.S. at 46 million acres (Ash et al., 2018). Together, these
137 crops cover more than 50% of cultivated land in the U.S.

138 Agricultural production is influenced by many factors other than land use, the most
139 important of which is weather. To control for the impact of weather on crop production, we
140 computed the average county-level temperature and precipitation for each day within a crop's
141 spatially varying growing season (Ramankutty, Evan, Monfreda, & Foley, 2008) from four-kilometer
142 gridded daily weather data provided by the PRISM Climate Group. From this daily data, we
143 computed three indicators of seasonal weather: growing degree days (GDDs), stress degree days
144 (SDDs), and total precipitation (TP). GDDs measure the accumulated degrees Celsius within a
145 crop-specific temperature range in which a crop's growth rate increases (Miller et al., 2001). The
146 tolerance range for corn and soy is 10–30° C and 0-30° C for winter wheat (Mesonet, 2017;
147 NDAWN, 2017). To model the negative effects of extreme temperature on crop production (Lobell
148 et al., 2013) we included SDDs, which are the total accumulated degrees Celsius above the maximum
149 GDD temperature threshold. To control for the impact of water availability on yields, we also
150 computed the TP or the cumulative sum of precipitation in millimeters throughout the growing
151 season.

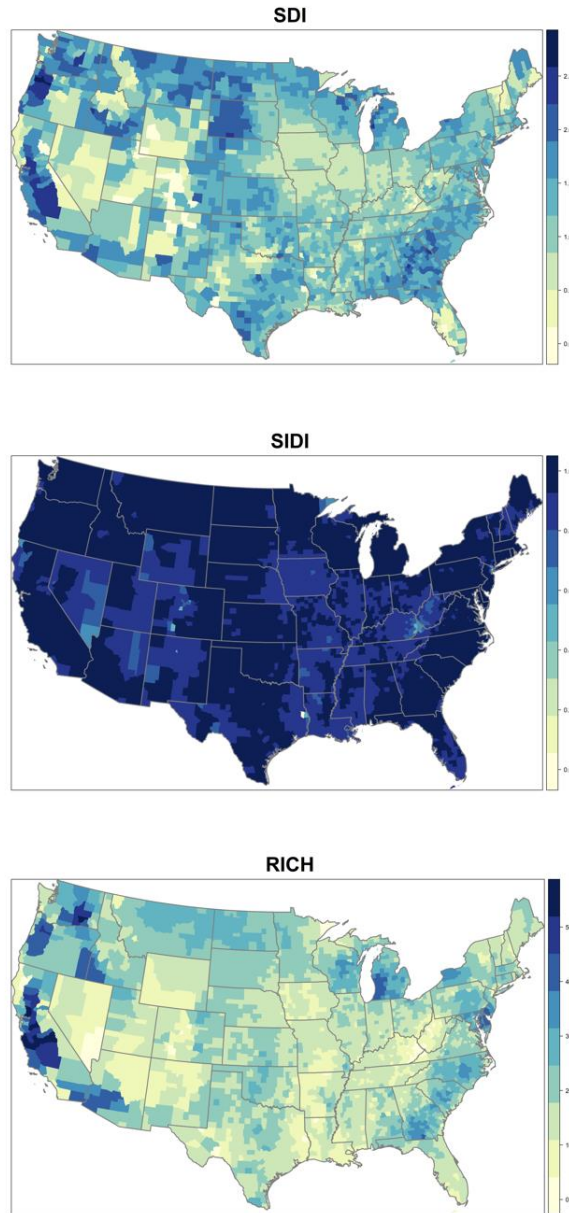
152 We used the USDA NASS Cropland Data Layer (CDL) as our indicator of land use. This
153 dataset classifies land use at a 30-meter resolution nationwide from 2008 to 2017 using satellite
154 imagery and extensive ground truth data. Using this data, we computed three county-scale
155 indicators of agricultural diversity: the Shannon Diversity Index, the Simpson Diversity Index, and
156 Richness. The Shannon Diversity Index (SDI) is a widely-used index of diversity that measures the
157 proportional abundance of each land use category in a given region (Aguilar et al., 2015; Gustafson,
158 1998; Turner, 1990). It incorporates both the number of land use categories and their relative
159 evenness on the landscape. The Simpson Diversity Index (SIDI) measures the probability that two
160 pixels selected at random belong to different land use categories. The SIDI gives more weight than
161 the SDI to common land use categories, i.e. rare land use categories will have a smaller effect on
162 SIDI than SDI. We also computed richness (RICH), or the number of unique land use categories in
163 a county. We extracted each index to the county-scale from the 100+ agricultural land use
164 categories included in the CDL (see the SI Appendix for the full list of categories). Each index varies
165 significantly across space, with the Midwestern U.S. generally exhibiting lower diversity than the
166 Southern and Western U.S. (Figure 2). By running our models with three commonly-used indices of
167 diversity, we can both test the sensitivity of our results to different operationalizations of diversity
168 and assess the extent to which different facets of diversity, e.g. abundance or relative extent, affect

169 yields. We include two additional spatially- and temporally-varying controls. The first is an indicator
 170 of the percent of irrigated land in a county extracted from the 250-meter gridded MiRAD dataset
 171 (Pervez & Brown, 2010). The second is an indicator of the prevalence and importance of a crop to
 172 a county's agricultural system, calculated as percent of agricultural acreage cultivated with the crop of
 173 interest.

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 175 **Table 1:** An overview of agricultural diversity indicators.

Index	Formula	Definition and advantages
Shannon Diversity Index (SDI)	$SDI = - \sum_{i=1}^k p_i \log(p_i)$	A measure of the abundance and evenness of land use categories. This index is sensitive to rare land use categories. Typical values are between 1.5 and 3.
Simpson Diversity Index (SIDI)	$SIDI = \frac{\sum n(n-1)}{N(N-1)}$	A measure of the abundance and evenness of land use categories. This index is <i>not</i> sensitive to rare land use categories. Values range from 0 to 1.
Richness (RICH)	Number of discrete land use types	A measure of the abundance of land use categories.

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 180 **Figure 2:** Variations in agricultural diversity as measured by the Shannon Diversity Index (SDI), the
 181 Simpson’s Diversity Index (SIDI), and Richness (RICH) for counties in the conterminous U.S. in
 182 2017.

183
 184 **2.2 Modeling**

185 The quantitative modeling in this study builds on work employing advanced statistical
 186 regression of cross-sectional time-series data—also known as panel data—to investigate known
 187 nonlinearities in the relationship between crop production and seasonal weather (Blanc & Schlenker,
 188 2017; Schlenker & Roberts, 2009). Agricultural production is strongly influenced by spatiotemporal
 189 context (Mendelsohn, & Massetti, 2017; Tack, Barkley, & Nalley, 2015); however, agro-climatic
 190 panel models typically employ frequentist statistics for which the incorporation of complex
 191 spatiotemporal dependency structures can be difficult and computationally expensive (Chatzopoulos
 192 & Lippert, 2015; Moore and Lobell, 2015). We leverage recent advances in Bayesian modeling

193 (Blangiardo & Cameletti, 2015; Mantovan & Secchi, 2010; Meehan & Gratton, 2016; Nelson &
 194 Burchfield, 2017) to control for the influence of spatiotemporal dependency in our estimation of the
 195 interactions between landscape, seasonal weather, and crop production. In addition to accounting
 196 for spatiotemporal dependencies which might otherwise bias our regression estimates, this approach
 197 has several advantages that are particularly relevant to our focus. First, it facilitates the estimation of
 198 known nonlinearities in the interactions between landscape, yield, and seasonal weather (Blanc &
 199 Schlenker, 2017; Butler & Huybers, 2015; Lobell et al., 2013; Schlenker & Roberts, 2009). Second, it
 200 controls for time-invariant spatially-varying factors (e.g. soil type, topography) and space-invariant
 201 temporally-varying factors (e.g. national policy changes, market variations) that influence yield,
 202 isolating yield variations driven by our variables of interest (Bivand, Gómez-Rubio, & Rue, 2015;
 203 Blangiardo & Cameletti, 2015; Meehan & Gratton, 2016). Third, this approach flexibly handles
 204 missing yield data by building model estimates using a combination of a specified likelihood
 205 function, specified prior probability distributions, and available data, providing multiple sources of
 206 information from which to build posterior effect estimates (Blangiardo & Cameletti, 2015).

207 We estimate a log-linear random-effects panel model that includes diversity (D), seasonal
 208 weather controls (GDD , SDD , TP), county-level spatial effects ($County$), and independent quadratic
 209 time trends for each region ($Time$) – identified using the Level III ecological regions provided by the
 210 US EPA (Figure 3) – to account for regionally-varying temporal changes that affect yield such as
 211 differences in technology adoption and management changes (Schlenker & Roberts, 2009).

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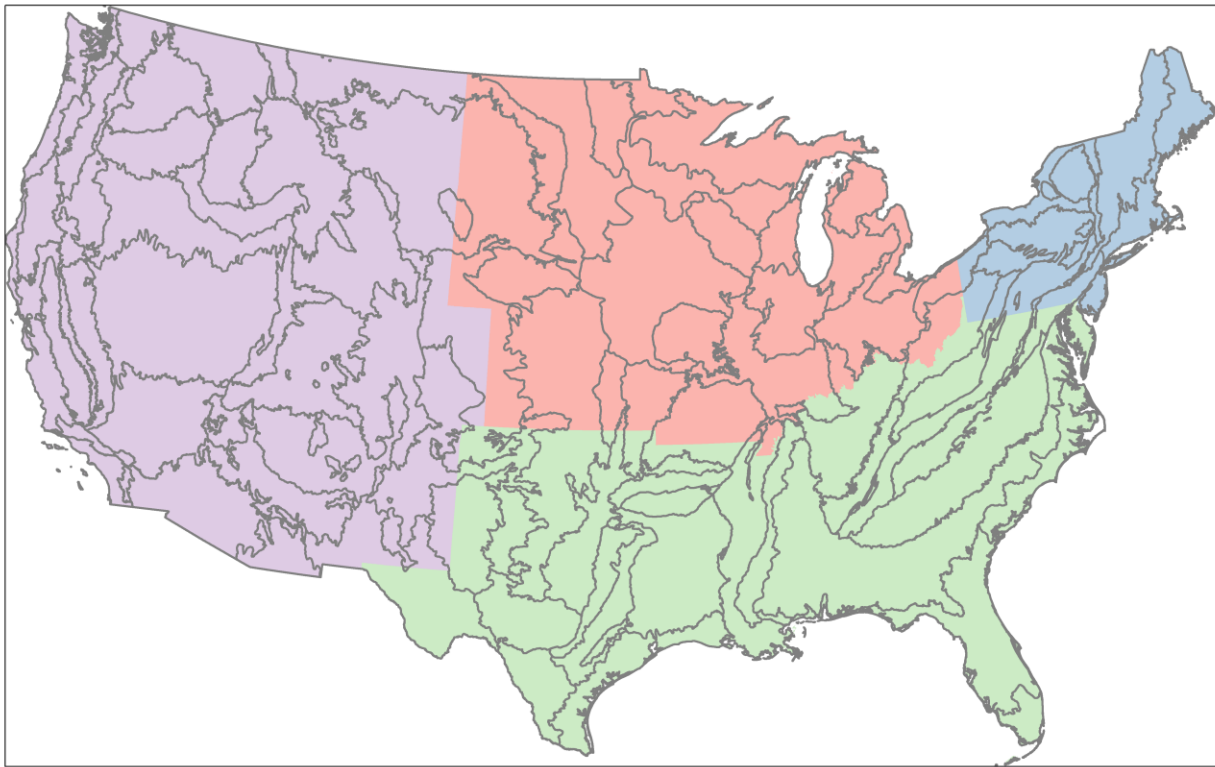
$$213 \log(Yield)_{ijt} = f(D)_{it} + f(TP)_{it} + f(GDD)_{it} + f(SDD)_{it} +$$

$$214 \quad \quad \quad Percent_Irrigated_i + Percent_Crop_{it} + f(County)_i + f(Time)_{jt}$$

215

216 where i indexes counties, j indexes regions, and t indexes year. County-level spatial effects ($County$)
 217 account for time-invariant factors associated with each county that influence yield including soil,
 218 topography, and non-dynamic sociocultural, infrastructure, and institutional factors. The county-
 219 level effects are modeled using a Besag-York-Mollié (BYM) structure which includes both
 220 exchangeable (iid) county random effects as well as conditional autoregressive structured (iCAR)
 221 residuals between counties. This formulation accounts for both random variation in yields across
 222 counties as well as spatial autocorrelation in yields across neighboring counties. $Percent_Irrigated$ and
 223 $Percent_Crop$ are linear controls that indicate the percent of irrigated land in a county and the percent
 224 of the county farmed with the crop of interest in each year, respectively. These controls account for
 225 known county characteristics that are expected to significantly impact yields. While most of the
 226 variance associated with these control variables is captured in the county-level spatial effects these
 227 variables are included as explicit controls in order to reduce chances of omitted variable bias (Blanc
 228 & Schlenker, 2017; Schlenker, Hanemann, & Fisher, 2007). Climate (TP , GDD , SDD) and diversity
 229 predictors (D , which includes SDI, SIDI, and RICH) are modeled using a first-order random-walk
 230 functional form. The random-walk structure allows the effect of these predictors to vary non-linearly
 231 (for example both low precipitation and very high precipitation tend to be associated with low
 232 productivity while moderate levels of precipitation tend to be associated with high productivity)
 233 while also considering that the effect of these predictors will be autocorrelated (e.g. similar values of
 234 TP will have a similar effect on yields).

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 237 **Figure 3:** Gray lines indicate the Level III ecological regions (US EPA, 2011) used for the quadratic
 238 time-trends. Colored areas represent the four major regions used in the regional models described in
 239 the Results and Discussion.

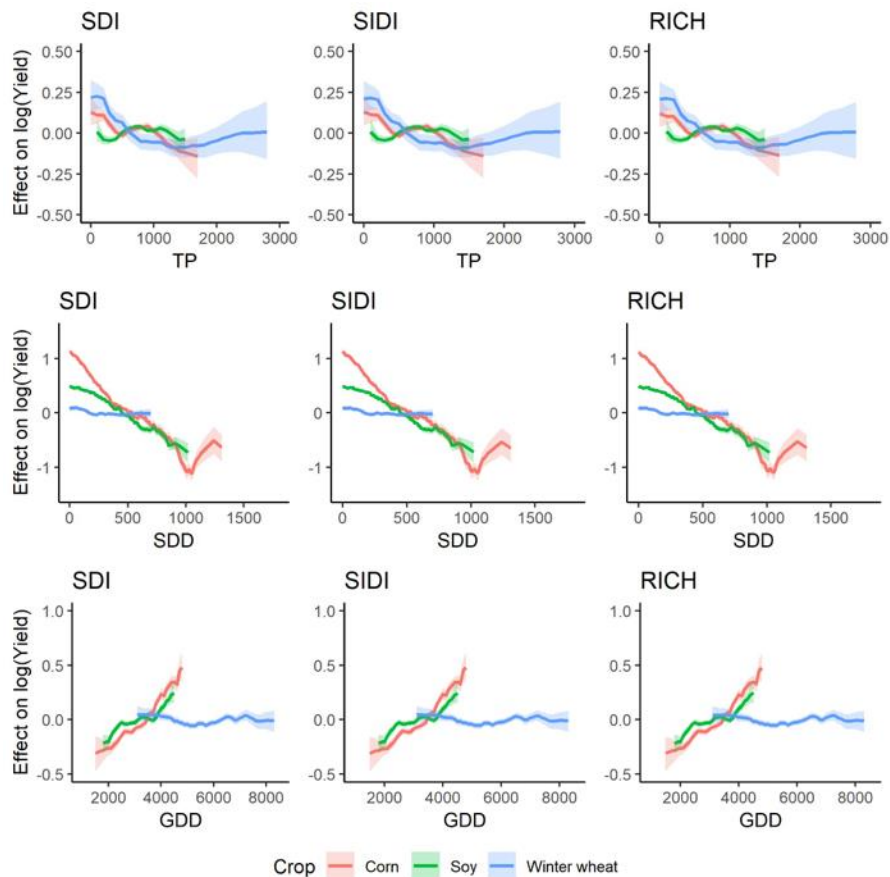
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 241 Our Bayesian models utilize a highly uninformative (reduced precision) prior distribution for
 242 linear effects and employ penalized complexity (PC) priors for the diversity, climate, and spatial
 243 effects (Simpson et al., 2017). The PC priors employ a scaling factor to specify priors based on
 244 sensible limits of the data (Simpson et al., 2017). We employed default and recommended settings
 245 for PC priors as provided by Simpson et al. (2017), yielding moderately informative priors. Model
 246 fit was evaluated using the deviance information criterion (DIC), the conditional predictive ordinate
 247 (CPO), the predictive probability integral transform (PIT), posterior predictive p-values, mean
 248 squared error (MSE) and Bayesian R-squared (R^2) (Blangiardo & Cameletti, 2015; Gelman,
 249 Goodrich, Gabry, & Ali, 2017; Gelman, A., Hill, J., 2007). Cross-validation of final models was
 250 conducted by re-estimating models with $\sim 86\%$ of the observations and comparing model
 251 predictions against the remaining held-out observations using MSE, R^2 , and the Nash-Sutcliffe
 252 Efficiency (NSE). In addition, model robustness checks were conducted to test sensitivity of results
 253 to the presence of control variables, data subsets, and prior specification (Schlenker, Hanemann, &
 254 Fisher, 2007). All models were estimated using the R-INLA package (Rue, Martino & Chopin, 2009)
 255 in R (R Core Team, 2014). Model scripts and additional information on model diagnostics and
 256 robustness checks are provided in the SI Appendix and on GitHub
 257 (https://github.com/eburchfield/Diversity_yield).

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Posterior Predictive Checks	MSE	0.0289	0.0289	0.0289	0.0174	0.0174	0.0174	0.0281	0.0281	0.0280
	R ²	0.7084	0.7082	0.7080	0.7521	0.7518	0.7522	0.7682	0.7687	0.7692
Cross-Validation	R ²	0.7397	0.7857	0.7112	0.7331	0.7736	0.7599	0.7362	0.7669	0.7940
	MSE	0.0423	0.0382	0.0413	0.0242	0.0258	0.0246	0.0391	0.0404	0.0404

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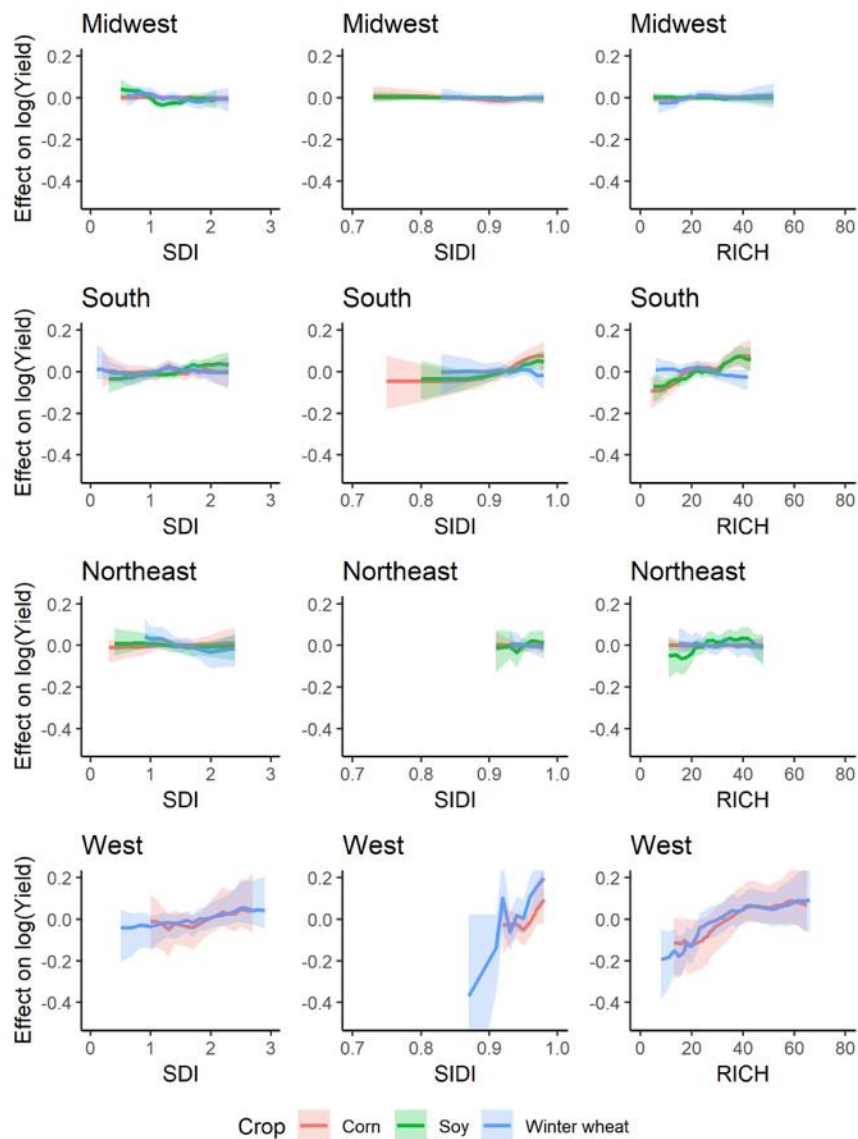
Our models also produce crop-specific response curves to seasonal weather (Figure 5). Estimated yield-weather interactions resemble those in the published literature, indicating that more GDDs increase yields, while higher SDDs decrease yields (Schlenker & Roberts, 2009). The idealized TP-yield curve is an inverse parabola (Rosenzweig et al., 2014), reflecting damages to crop production under extremely low and high precipitation conditions. Corn and soy exhibit this response, while winter wheat yields increase only at high levels of precipitation. This may be attributable to the fact that unlike corn and soy, which are grown over the summer and harvested in early fall, winter wheat is planted in the fall and is harvested for grain the following spring. For all crops, very low seasonal precipitation is associated with higher yields. This may be due to the relatively short time-frame of our panel (2010-2016) as well as the importance of irrigation as a buffer against low precipitation over this period.



301

302 **Figure 5:** The response of yields to changes in total seasonal precipitation in millimeters (TP),
 303 growing degree days (GDDs), and stress degree days (SDDs) in degrees Celsius.

304
 305 While the national models *control* for regional differences, they do not explicitly *model* the
 306 ways in which these differences influence diversity-production interactions. To assess how the
 307 relationship between agricultural diversity and crop production varies across space, we re-estimated
 308 models in four major regions of the U.S.: the South, Northeast, Midwest, and Western U.S. (Figure
 309 3). The models suggest that in places where large-scale farming is less common for edaphic,
 310 topographic, cultural, or infrastructural reasons—such as the Southern and Western U.S.—
 311 agricultural diversity has a far greater impact on crop production (Figure 6). Midwestern and
 312 Northeastern agricultural systems are relatively insensitive to diversification, while Western systems
 313 show strong and sustained yield responses to all indicators of diversity. The Southern U.S. shows the
 314 most variability across crops, with positive yield responses for corn and soy, and slightly negative
 315 yield responses for winter wheat.



317
 318 **Figure 6:** Regional differences in the effect of agricultural diversification on crop production.

319

320 4.0 DISCUSSION

321

322 Our results suggest agricultural diversification can directly benefit agricultural systems.
323 Yields of corn and winter wheat increase by as much as 20% in highly diversified agricultural
324 systems, and soy yields increase by nearly 5%. Our findings also indicate that (1) crop production is
325 more responsive to the *number* of agricultural land use categories in a region than the relative
326 cultivated *extent* of each category and that (2) increasing agricultural diversity in regions that are
327 already diverse brings the highest yield gains. These results provide strong empirical support for *why*
328 we should consider agricultural diversification. In what follows, we discuss how these models can
329 also give us a better sense of where, when, and how to diversify.

330

331 4.1 Where to diversify? The importance of regional variability.

332 We find that agricultural diversification has a stronger impact on corn and winter wheat than
333 soy nationally, but these effects vary across regions (Figure 6). For example, winter wheat shows
334 markedly different responses to increased agricultural diversity in the Western and Southern U.S.,
335 while soy – relatively unresponsive to diversification in the national models – shows significant
336 responses to diversification in the Southern and Northeastern U.S. The regional models highlight
337 two important findings. First, differences in diversity-productivity curves across crops and indices as
338 seen in the national models are less significant than differences in diversity-productivity curves
339 across regions. This suggests that regional factors may play a larger role in moderating the diversity-
340 productivity relationship than crop- and index-specific factors. Second, the regional models
341 correspond with published literature indicating that the ways in which diversity interacts with crop
342 production varies significantly across agricultural, climatic, ecological, and socio-cultural contexts
343 (Balvanera et al., 2006; Loreau et al., 2001; Swift et al., 2004; Tilman et al., 2014; Zak et al., 2003).
344 The spatial variability of our findings highlights the fundamental challenge of *scale* in agro-ecological
345 research. Large scale models, such as those presented in this paper, provide empirical support for
346 interventions that may sustainably increase agricultural productivity, but are limited in their ability to
347 provide context-specific recommendations to support agricultural decision-making. Conversely,
348 field-scale analyses can provide specific recommendations for farmers but are limited in their
349 generalizability across regions.

350 Despite regional variability, in very few cases does diversification *decrease* yields. A shift from
351 low to moderate levels of agricultural diversity decreases yields for winter wheat (national model);
352 however, this is only the case for SDI, suggesting that how land uses are partitioned, as opposed to
353 the number of land uses itself, is driving this effect. Increasing diversification is also associated with
354 decreased winter wheat yields in the Northeastern and Southern regional models and with decreased
355 soy yields in the Midwestern regional model, however these decreases are not significant.

356

357 4.2 When to diversify? The importance of contextual variability.

358 Our primary objective in modeling yield response to multiple indicators of diversity (SDI,
359 SIDI, RICH) is to test model sensitivity to operationalizations of diversity; however, model
360 differences also provide insights into how agricultural systems respond to different facets of
361 diversity. Our results indicate that the yields of corn and winter wheat are more responsive to SDI
362 and RICH (Figure 4A and 4C) than SIDI (Figure 4B). The SIDI is less sensitive than SDI and RICH
363 to rare land use categories, meaning that a small increase in agricultural diversity in a system
364 dominated by a single crop will increase both the SDI and RICH much more than the SIDI.
365 Therefore, the relative responsiveness of corn and winter wheat yields to changes in SDI and RICH
366 suggests that these crops are more sensitive to the *number* of distinct crops in a county rather their

367 relative cultivated extent. This finding merits further exploration, as it indicates that cultivating small
368 areas of a landscape with a new crop could increase agricultural productivity.

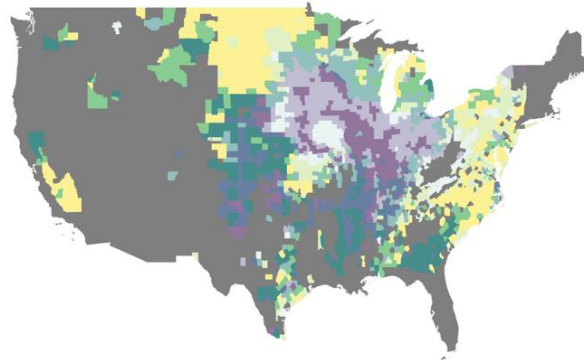
369 By estimating non-linearities in the diversity-productivity relationship, we can also identify
370 the specific ranges of agricultural diversity that have the highest potential impact on crop
371 production. The linear response of yields to RICH indicates that adding a new crop to an
372 agricultural system operating at any level of diversity can increase yields; however, the shape of the
373 SDI and SIDI curves in Figure 4 suggests that increasing agricultural diversity in systems that are
374 already diverse brings the highest yield gains. For example, increasing SDI from 2 to 3 increases
375 yields of corn and winter wheat by approximately 10 and 20% respectively. Similarly, increasing
376 SIDI from 0.9 to 1.0 increases yields of corn and winter wheat by nearly 10%. Yield gains for corn
377 and soy are much lower when systems move from low to moderate agricultural diversity. In the case
378 of winter wheat, diversification in this range may actually decrease yields. We hypothesize that this
379 is, in part, due to heavy reliance on external mechanized and chemical inputs in specialized
380 monoculture systems that offset (at least in the short-term) the negative impacts of diversity loss.

381 Increasing agricultural diversity in regions that are already diverse has positive effects on
382 yields of all crops across indicators of diversity and across regions. There is far more variability in
383 crop response to diversification in systems with low agricultural diversity. These findings emerge
384 both at the national and regional scales, with the Midwestern U.S.—a region dominated by
385 monoculture systems—showing weaker yield responses to agricultural diversification than other
386 regions of the U.S. This illustrates the importance not only of the regional variability discussed
387 above, but of contextual variability, or the impact of current landscape composition on the
388 effectiveness of diversification.

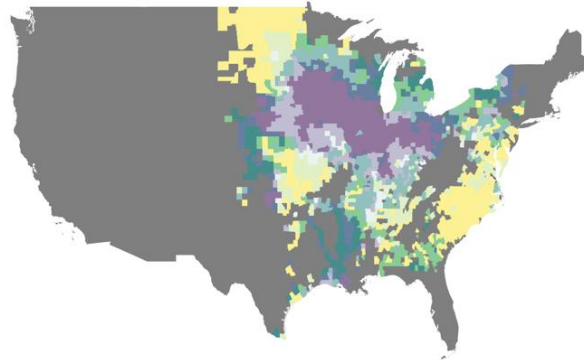
389 Figure 7 classifies systems in terms of their combined diversity and productivity. Diverse and
390 productive systems are shown in dark green, while simplified and productive systems—largely
391 concentrated in the Midwestern U.S.—are shown in dark purple. This figure highlights the
392 importance of regional variability in diversity-productivity interactions but can also help to target
393 regions where agricultural diversification may have the highest impact. Given our finding that
394 agricultural diversification has the highest impact in systems that are already fairly diverse,
395 diversification efforts targeted in regions of low to moderate agricultural productivity and moderate
396 to high agricultural diversity (light greens and yellow regions) may have the highest impact.

397

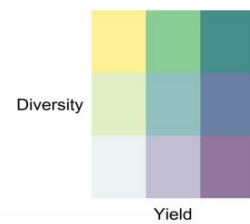
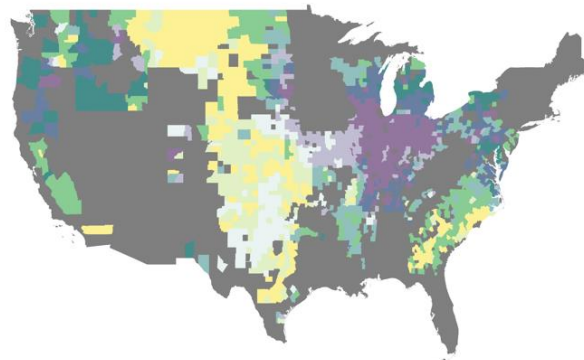
Corn



Soy



Winter wheat



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Figure 7: Bivariate choropleth constructed by binning county-level spatial effects and SDI into thirds. We use the county-level spatial effects from the model described in Section 2.2 run without diversity predictors as our indicator of yields. These effects capture the average yield in a county given the non-diversity predictors in our model (seasonal weather, irrigation, and acreage). Regions in dark green are both highly diverse and highly productive. Yellow regions are highly diverse, but low productivity, and purple regions are highly productive but low diversity.

406 Why might simple agricultural systems exhibit a more varied response to diversification than
407 diverse agricultural systems? Simple agricultural systems, such as the monoculture systems that
408 dominate much of the Midwestern U.S., tend to be highly specialized, intensively managed, and
409 heavily reliant on external petro-chemical and mechanical inputs (Altieri, 1999; Foley et al., 2011;
410 Kremen, Iles, & Bacon, 2012). These systems have some of the highest yields on the planet (USDA-
411 FAS, 2017); however, these yields are not without environmental consequence (Rabalais et al., 2002;
412 Kremen & Miles, 2012). We hypothesize that benefits from diversification in these systems are
413 drowned out by the yield gains brought by intensive management. We note that, except in the case
414 of winter wheat in a subset of models, crop production does not *decrease* with diversification. In fact,
415 yield responses to RICH are near-linear in national models and consistently positive in the regional
416 models across all crops. While we acknowledge the significant cost and barriers to diversification in
417 monoculture systems (Blesh & Wolf, 2014; Roesch-McNally et al., 2018; Lin, 2011; Roesch-
418 McNally, Arbuckle, & Tyndall, 2018), this result suggests that simple interventions, such as adding a
419 small area cultivated with a new crop, could significantly increase crop production even in the most
420 simplified systems.

421

422 **4.3 How to diversify? A landscape “commons”**

423 This investigation of the relationship between agricultural diversity and crop productivity has
424 important implications for farmers and land managers across the U.S. Our results suggest that by
425 increasing the compositional heterogeneity of crops within a landscape, farmers can significantly
426 increase yields. Furthermore, our models suggest that it is the *number* of crops cultivated rather than
427 their cultivated extent that can bring greater yield benefits. This suggests that relatively simple
428 interventions such as adding a new crop cultivated on a small extent could increase agricultural
429 system productivity. This also implies that a single farmer does not necessarily need to abandon
430 monoculture to see yield gains; conversely, a diversified farmer may not see gains in productivity
431 when cultivating in a simplified landscape. These dynamics emphasize the importance of conducting
432 analyses at a *landscape* scale and of re-conceptualizing working landscapes, as well as the ecosystem
433 services they generate, as common pool resources (Ostrom, 1990; Zhang et al., 2007). The benefits
434 of agricultural diversification flow across property boundaries and associated costs may not be fairly
435 spread across users. There is a growing need to understand land use diversification beyond
436 individual farmer decisions and within the feasibility of coordinated landscape management and
437 connectivity (DeClerck, Estrada-Carmona, Garbach, & Martinez-Salinas, 2015).

438

439 **5.0 CONCLUSION**

440

441 Human-induced reductions in diversity have had negative impacts on ecosystem function
442 comparable to those from elevated carbon dioxide concentrations, nitrogen deposition, fire, and
443 drought (Hooper et al., 2012; Tilman, Reich, & Isbell, 2012). Agriculture is a significant driver of this
444 diversity loss and will likely remain as such if current practices persist (Bommarco, Kleijn, & Potts,
445 2013; Hendrickx et al., 2007; Landis, 2017; McDaniel, Tiemann, & Grandy, 2014; Tiemann, Grandy,
446 Atkinson, Marin-Spiotta, & McDaniel, 2015; Tscharrntke et al., 2012). In this paper, we assess
447 whether and how increasing agricultural diversity affects agricultural health and productivity. Our
448 models provide strong evidence at national and regional scales that agricultural diversification—an
449 intervention with known ecosystem benefits—can increase crop production. This suggests that
450 agricultural diversification could serve as a key land use strategy to boost agricultural production
451 while preserving ecosystem function and integrity. These findings are relatively consistent across
452 crops, indices of diversity, and regions of the U.S. However, these findings do not identify the

453 specific causal mechanisms underlying the relationship between landscape diversity and crop
454 production. A limitation of this study is the inability to account for local-scale factors and sub-
455 annual variability such as application of fertilizer and pesticides for which data availability is limited
456 and natural disaster events (see Figure S1). The regional variability in our models, highlights the
457 continued importance of local- and meso-scale analyses to assess the complex assemblage of socio-
458 ecological factors that mediate the diversity-productivity relationship across space and time. In
459 addition, the time frame for which the USDA Cropland Data Layer land use information is available
460 limits this study to a relatively short window of time, making these model results more sensitive to
461 annual variation as shown in Figures S2-S4. Additional research is needed to identify the social and
462 ecological moderators of the diversity-productivity relationship and key barriers to diversification
463 such as capital and cost, risk perceptions and behavior, market dynamics, institutional constraints,
464 and changing climate (Burchfield & Poterie, 2018; Di Falco & Perrings, 2005; Roesch-McNally,
465 Arbuckle, & Tyndall, 2018). Our hope is that the empirical evidence provided in this paper will
466 motivate future initiatives to identify these barriers and moderators.

467

468 **References**

469

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